## Ordinary Least Squares, continued

EDS 222

Tamma Carleton
Fall 2023

## Announcements/check-in

- Assignment \#1: Grades posted
- Please ensure your .html file is compiled and pushed to GitHub
- Please do not push data to GitHub (generally a good rule to follow)
- Sandy to go over some areas of confusion


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- Assignment \#2: Due 10/20, 5pm


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- Assignment \#2: Due 10/20, 5pm
- Reiteration of COVID/illness policy


## Today

Notes on OLS

- Outliers, missing data


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Measures of model fit

- Coefficient of variation $R^{2}$


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- In R, interpretation


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## Categorical variables

- In R, interpretation

Multiple linear regression

- Adding independent variables, interpretation of results

Notes on OLS

## Outliers

Because OLS minimizes the sum of the squared errors, outliers can play a large role in our estimates.

## Common responses

- Remove the outliers from the dataset
- Replace outliers with the $99^{\text {th }}$ percentile of their variable (winsorize)
- Take the $\log$ of the variable (This lowers the leverage of large values -- why?)
- Do nothing. Outliers are not always bad. Some people are "far" from the average. It may not make sense to try to change this variation.


## Missing data

Similarly, missing data can affect your results.
R doesn't know how to deal with a missing observation.
$1+2+3+N A+5$
\#> [1] NA
If you run a regression ${ }^{\dagger}$ with missing values, R drops the observations missing those values.
If the observations are missing in a nonrandom way, a random sample may end up nonrandom.

Measures of model fit

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Goal: quantify how "well" your regression model fits the data
General idea: Larger variance in residuals suggests our model isn't very predictive


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## Coefficient of determination

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- Interpretation of $R^{2}$ : share of the variance in $y$ that is explained by your regression model

$$
\begin{gathered}
S S R=\text { sum of squared residuals }=\sum_{i}\left(y_{i}-\hat{y}_{i}\right)^{2}=\sum_{i} e_{i}^{2} \\
S S T=\text { total sum of squares }=\sum_{i}\left(y_{i}-\bar{y}\right)^{2} \\
R^{2}=1-\frac{S S R}{S S T}=1-\frac{\sum_{i} e_{i}^{2}}{\sum_{i}\left(y_{i}-\bar{y}\right)^{2}}
\end{gathered}
$$

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- $R^{2}$ varies between 0 and 1: Perfect model with $e_{i}=0$ for all $i$ has $R^{2}=1 . R^{2}=0$ if we just guess the mean $\bar{y}$.
- In more complex models, $R^{2}$ is not the same as the square of the correlation coefficient. You should think of them as related but distinct concepts.


## Coefficient of determination

About 49\% of the variation in ozone can be explained with temperature alone!

```
#>
#> Call:
#> lm(formula = Ozone ~ Temp, data = airquality)
#>
#> Residuals:
\begin{tabular}{lrrrrr} 
\#> & Min & \(1 Q\) & Median & 3Q & Max \\
\#> & -40.729 & -17.409 & -0.587 & 11.306 & 118.271
\end{tabular}
#>
#> Coefficients:
#> Estimate Std. Error t value Pr(>|t|)
#> (Intercept) -146.9955 18.2872 -8.038 9.37e-13 ***
#> Temp 2.4287 0.2331 10.418<2e-16 ***
#> ---
#> Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
#>
#> Residual standard error: 23.71 on 114 degrees of freedom
#> (37 observations deleted due to missingness)
#次Multiple R-squared: 0.4877, Adjusted R-squared: 0.4832
#> F-statistic: 108.5 on 1 and 114 DF, p-value: < 2.2e-16
```


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Definition: \% of variance in $y$ that is explained by $x$ (and any other independent variables)

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- Higher $R^{2}$ does not mean a model is "better" or more appropriate
- Predictive power is not often the goal of regression analysis (e.g., you may just care about getting $\beta_{1}$ right)
- If you are focused on predictive power, many other measures of fit can be appropriate (to discuss in machine learning)
- Always look at your data and residuals!


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- Predictive power is not often the goal of regression analysis (e.g., you may just care about getting $\beta_{1}$ right)
- If you are focused on predictive power, many other measures of fit can be appropriate (to discuss in machine learning)
- Always look at your data and residuals!
- Like OLS in general, $R^{2}$ is very sensitive to outliers. Again...always look at your data!


## Coefficient of determination

Here, $R^{2}=0.94$ for a model of $y=\beta_{0}+\beta_{1} x+\epsilon$. Does that mean a linear relationship with $x$ is appropriate?


## Coefficient of determination

Here, $R^{2}=0$ for a model of $y=\beta_{0}+\beta_{1} x+\epsilon$. Does that mean there is no relationship between these variables?


Indicator/categorical variables

## Categorical variables

We have been talking a lot about numerical variables in linear regression...

- Ozone levels
- Crab size
- Temperature and precipitation amounts
- etc.


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- Presence/absence of a species
- In/out of compliance with a pollution standard
- etc.


## Categorical variables

We have been talking a lot about numerical variables in linear regression...

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...but a lot of variables of interest are categorical:
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- Presence/absence of a species
- In/out of compliance with a pollution standard
- etc.

How do we execute and interpret linear regression with categorical data?

## Categorical variables

We use dummy or indicator variables in linear regression to capture the influence of a categorical independent variable ( $x$ ) on a continuous dependent variable ( $y$ ).

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For example, let $x$ be a categorical variable indicating the gender of an individual. Suppose we are interested in the "gender wage gap", so y is income We estimate:

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y_{i}=\beta_{0}+\beta_{1} M A L E_{i}+\varepsilon_{i}
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Interpretation [draw it]:

- $M A L E_{i}$ is an indicator variable that $=1$ when $i$ is male ( 0 otherwise)
- $\beta_{0}=$ average wages if $i$ is not male
- $\beta_{0}+\beta_{1}=$ average wages if $i$ is male
- $\beta_{1}=$ average difference in wages between males and females


## Categorical variables

For a categorical variable with two "levels", the OLS slope coefficient is the difference in means across the two groups


## Categorical variables

What if I have many categories?

- E.g., species, education level, age group, ...

For example, let $x$ be a categorical variable indicating the species of penguin, and $y$ is body mass. We estimate:

$$
y_{i}=\beta_{0}+\beta_{1} S P E C I E S_{i}+\varepsilon_{i}
$$

Where species can be one of:

- Adelie
- Chinstrap
- Gentoo


## Categorical variables

```
library(palmerpenguins)
head(penguins)
#> # A tibble: 6 x 8
#> species island bill_length_mm bill_depth_mm flipper_l...1 body_... }\mp@subsup{}{}{2}\mathrm{ sex year
#> <fct> <fct> <dbl> <dbl> <int> <int> <fct> <int>
#> 1 Adelie Torgersen 39.1 18.7 181 3750 male 2007
#> 2 Adelie Torgersen 39.5 17.4 186 3800 fema... 2007
#> 3 Adelie Torgersen 40.3 18 195 3250 fema... 2007
#> 4 Adelie Torgersen NA NA NA NA <NA> 2007
#> 5 Adelie Torgersen 36.7 19.3 193 3450 fema... 2007
#> 6 Adelie Torgersen 39.3 20.6 190 3650 male 2007
#> # ... with abbreviated variable names 1flipper_length_mm, 2body_mass_g
class(penguins$species)
#> [1] "factor"
```


## Categorical variables

```
summary(lm(body_mass_g ~ species, data = penguins))
#>
#> Call:
#> lm(formula = body_mass_g ~ species, data = penguins)
#>
#> Residuals:
\begin{tabular}{lrrrrr} 
\#> & Min & 1Q & Median & 3Q & Max \\
\#> & -1126.02 & -333.09 & -33.09 & 316.91 & 1223.98
\end{tabular}
#>
#> Coefficients:
#> Estimate Std. Error t value Pr(>|t|)
#> (Intercept) 3700.66 37.62 98.37 <2e-16 ***
#> speciesChinstrap 32.43 67.51 0.48 0.631
#> speciesGentoo 1375.35 56.15 24.50 <2e-16 ***
#> _-_
#> Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
#>
#> Residual standard error: 462.3 on 339 degrees of freedom
#> (2 observations deleted due to missingness)
#> Multiple R-squared: 0.6697, Adjusted R-squared: 0.6677
#> F-statistic: 343.6 on 2 and 339 DF, p-value: < 2.2e-16
```


## Categorical variables

What is going on here?? One $x$ variable turned into multiple slope coefficients? :

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$R$ is turning our regression

$$
y_{i}=\beta_{0}+\beta_{1} S P E C I E S_{i}+\varepsilon_{i}
$$

where SPECIES is a categorical variable indicating one of three species, into:

$$
y_{i}=\beta_{0}+\beta_{1} C H I N S T R A P_{i}+\beta_{2} \text { GENTOO }_{i}+\varepsilon_{i}
$$

where CHINSTRAP and GENTOO are dummy variables for the Chinstrap and Gentoo species, respectively.

## Categorical variables

When your categorical variable takes on $k$ values, R will create dummy variables for $k-1$ values, leaving one as the reference group:
\#> Coefficients:

| \#> | Estimate | Error | value | $>\|t\|)$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| \#> (Intercept) | 3700.66 | 37.62 | 98.37 | $<2 e-16$ | *** |
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\#> Coefficients:


To evaluate the outcome for the reference group, set the dummy variables equal to zero for all other groups.
Q: What is the average body mass of an Adelie species?
Q: What is the difference in body mass between Chinstrap and Adelie?

## Multiple linear regression

## More explanatory variables

We're moving from simple linear regression (one outcome variable and one explanatory variable)

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y_{i}=\beta_{0}+\beta_{1} x_{i}+u_{i}
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Why? We can better explain the variation in $y$, improve predictions, avoid omitted-variable bias (i.e., second assumption needed for unbiased OLS estimates), ...

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We will dig into each of these here, and you will see these questions in other MEDS courses

## Multiple regression

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## Multiple regression

The intercept and categorical variable $x_{2}$ control for the groups' means.


## Multiple regression

$\hat{\beta}_{1}$ estimates the relationship between $y$ and $x_{1}$ after controlling for $x_{2}$. This is often called the "parallel slopes" model (one slope $\beta_{1}$ for each of the groups in $x_{2}$ )


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## Multiple regression

With many explanatory variables, we visualizing relationships means thinking about hyperplanes

$$
y_{i}=\beta_{0}+\beta_{1} x_{1 i}+\beta_{2} x_{2 i}+\ldots+\beta_{k} x_{k i}+u_{i}
$$

Math notation looks very similar to simple linear regression, but conceptually and visually multiple regression is very different

## Multiple regression

Interpretation of coefficients

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- E.g., how much do wages increase with one more year of education, holding gender fixed?


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- This is an "all else equal" interpretation
- E.g., how much do wages increase with one more year of education, holding gender fixed?
- E.g., how much does ozone increase when temperature rises, holding NOx emissions fixed?


## Tradeoffs

There are tradeoffs to consider as we add/remove variables:

## Fewer variables

- Generally explain less variation in $y$
- Provide simple interpretations and visualizations (parsimonious)
- May need to worry about omitted-variable bias


## More variables

- More likely to find spurious relationships (statistically significant due to chance-does not reflect a true, population-level relationship)
- More difficult to interpret the model
- You may still miss important variables-still omitted-variable bias


## Omitted-variable bias

You will study this in much more depth in EDS 241, but here's a primer.
Omitted-variable bias (OVB) arises when we omit a variable that

1. affects our outcome variable $y$
2. correlates with an explanatory variable $x_{j}$

As it's name suggests, this situation leads to bias in our estimate of $\beta_{j}$. In particular, it violates Assumption 2 of OLS from last week.

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Note: OVB Is not exclusive to multiple linear regression, but it does require multiple variables affect $y$.

## Omitted-variable bias

## Example

Let's imagine a simple model for the cancer rates in census tract $i$ :

$$
\text { Cancer rate }_{i}=\beta_{0}+\beta_{1} \mathrm{UV} \text { radiation }_{i}+\beta_{2} \mathrm{TRI}_{i}+u_{i}
$$

where

- UV radiation ${ }_{i}$ gives the average UV radiation in tract $i(\mathrm{~mW} / \mathrm{cm} \$ \wedge 2 \$)$
- $\mathrm{TRI}_{i}$ denotes an indicator variable for whether tract $i$ has a Toxics Release Inventory facility
thus
- $\beta_{1}$ : the change in cancer rate associated with a $1 \mathrm{~mW} / \mathrm{cm} \$ \wedge 2 \$$ increase in UV radiation (ceteris paribus)
- $\beta_{2}$ : the difference in avg. cancer rates between TRI and non-TRI census tracts (ceteris paribus) If $\beta_{2}>0$, then TRI tracts have higher cancer rates


## Omitted-variable bias

"True" relationship: Cancer rate ${ }_{i}=20+0.5 \times \mathrm{UV}$ radiation $_{i}+10 \times \mathrm{TRI}_{i}+u_{i}$
The relationship between cancer rates and UV radiations:


## Omitted-variable bias

Biased regression estimate: Cancer $\widehat{r a t e}_{i}=31.3+-0.9 \times$ UV radiation $_{i}$


## Omitted-variable bias

Recalling the omitted variable: TRI (non-TRI and TRI)


## Omitted-variable bias

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## Omitted-variable bias

Unbiased regression estimate: $\widehat{\text { Cancer rate }}{ }_{i}=20.9+0.4 \times$ UV radiation $_{i}+9.1 \times \mathrm{TRI}_{i}$


Slides created via the R package xaringan.
Some slide components come from Ed Rubin's awesome course materials.

